

ON THE ESTIMATION OF MEAN WHEN POPULATION VARIANCE IS KNOWN

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SUMMARY

A class of estimators for population mean when the population variance is known has been suggested. It is shown that the optimum estimator in the class has a smaller bias as compared to estimators proposed earlier.

Keywords : Optimum estimator, mean square error, large sample approximations.

Introduction

Let y_1, y_2, \dots, y_n be a random sample of size n from a population with unknown mean \bar{Y} and known variance σ^2 . Upadhyaya and Srivastava [2] proposed the estimator

$$\bar{Y}_1 = \bar{y} + \sigma^2/n\bar{y} \quad (1.1)$$

and obtained the bias and mean squared error (MSE) to order $O(n^{-2})$ as

$$\text{Bias}(\bar{Y}_1) = (\sigma^2/n\bar{Y})(1 + \sigma^2/n\bar{Y}^2) \quad (1.2)$$

$$\text{and } \text{MSE}(\bar{y}_1) = (\sigma^2/n) (1 - \sigma^2/n\bar{Y}^2) \quad (1.3)$$

Upadhyaya and Singh [1] proposed an equally efficient estimator

$$\bar{y}_2 = \bar{y} + \sigma^2 \bar{Y}/n\bar{y}(\bar{y}^2 + \sigma^2) \quad (1.4)$$

but with smaller bias to order $O(n^{-2})$ as

$$\text{Bias}(\bar{y}_2) = \sigma^2/n\bar{Y} \quad (1.5)$$

The purpose of this paper is to consider a class of estimators of the mean when population variance is known. The large sample approximations have also been investigated.

2. Proposed Class of Estimators

Assuming the population variance σ^2 to be known, we consider the class of estimators for the mean \bar{Y} as

$$\bar{y}_3 = \bar{y} + K\sigma^2 \bar{y}\delta - \sigma^2/n\bar{y} \quad (2.1)$$

where $\delta = (n\bar{y}^2 + \sigma^2)^{-1}$ and K is the characterizing scalar to be chosen suitably.

The bias, MSE and K_{opt} for minimum MSE are

$$\text{Bias}(\bar{y}_3) = K\sigma^2 E(\bar{y}\delta) - \sigma^2 E(n\bar{y})^{-1} \quad (2.2)$$

$$\begin{aligned} \text{MSE}(\bar{y}_3) &= E(\bar{y}_3 - \bar{Y})^2 + K^2 \sigma^4 E(\bar{y}\delta)^2 + \sigma^4 E(n\bar{y})^{-2} \\ &\quad + 2K\sigma^2 E(\bar{Y}^2\delta) - 2K\sigma^2 \bar{Y} E(\bar{y}\delta) \\ &\quad - 2K\sigma^4 n^{-1} E(\delta) + 2\sigma^2 \bar{Y} E(n\bar{y})^{-1} - 2\sigma^2 n^{-1} \end{aligned} \quad (2.3)$$

$$\text{and } K_{\text{opt}} = [\bar{Y} E(\bar{y}\delta) + \sigma^2 n^{-1} E(\delta) - E\bar{y}^2\delta] / [\sigma^2 E(\bar{y}\delta)^2] \quad (2.4)$$

The expectations involved in these expressions are mathematically intractable. We, therefore, resort to derive the large sample approximations.

3. Large Sample Approximations

We write

$$\bar{y} = \bar{Y} + \varepsilon$$

where ϵ is of the order $O(n^{-1/2})$ with $E(\epsilon) = 0$ and $E(\epsilon^2) = \sigma^2/n$. Choosing n large enough so that $|\epsilon|/Y < 1$, we may easily get the expectations by retaining terms of order $O(n^{-2})$ as

$$\begin{aligned} E(ny)^{-1} &= (1 + \sigma^2/n\bar{Y}^2)/n\bar{Y}, E(n^2\bar{y}^2)^{-1} = (n^2\bar{Y}^2)^{-1} \\ E(\delta) &= (1 + 2\sigma^2/n\bar{Y}^2)/n\bar{Y}^2, E(y\delta) = (n\bar{Y})^{-1} \end{aligned} \quad (3.1)$$

$$\text{and } E(y\delta)^2 = (n^2\bar{Y}^2)^{-1} \text{ and } E(y^2\delta) = 1 - \sigma^2/n\bar{Y}^2/n$$

From (2.2), (2.3), (2.4) and (3.1) we get (upto terms of order $O(n^{-2})$)

$$\text{Bias}(\bar{y}_2) = \sigma^2(K-1 - \sigma^2/n\bar{Y}^2)/n\bar{Y} \quad (3.2)$$

$$\text{MSE}(\bar{y}_2) = \sigma^2[1 + (K-1)(K-3)\sigma^2/n\bar{Y}^2]/n \quad (3.3)$$

$$\text{and } K_{\text{opt.}} = 2. \quad (3.4)$$

4. Comparison

For $K_{\text{opt.}}$, the estimator \bar{y}_2 , its bias and MSE become

$$\bar{y}_2 = y + \sigma^2 y (1 - \sigma^2/n\bar{Y}^2)/(n\bar{y}^2 + \sigma^2) \quad (4.1)$$

$$\text{Bias}(\bar{y}_2) = (\sigma^2/n\bar{Y}) (1 - \sigma^2/n\bar{Y}^2) \quad (4.2)$$

$$\text{and } \text{MES}(\bar{y}_2) = (\sigma^2/n) (1 - \sigma^2/n\bar{Y}^2) \quad (4.3)$$

From (1.2), (1.5) and (4.2), we have $\text{Bias}(\bar{y}_2) < \text{Bias}(\bar{Y}_2) < \text{Bias}(\bar{Y}_1)$ upto $O(n^{-2})$. And all three estimators have equal MSE. Therefore, the estimator \bar{y}_2 is an improvement over \bar{y}_1 and \bar{Y}_1 .

REFEREECES

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